

# Finite Memory Estimation of Infinite Memory Processes

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**Abstract**—Stationary ergodic processes with finite alphabets are approximated by finite memory processes based on an  $n$ -length realization of the process. Under the assumptions of summable continuity rate and non-nullness, a rate of convergence in  $\bar{d}$ -distance is obtained, with explicit constants. Asymptotically, as  $n \rightarrow \infty$ , the result is near the optimum.

## I. INTRODUCTION

This paper deals with estimation of stationary ergodic processes based on a sample, an observed finite realization of the process, of length  $n$ . We consider the  $\bar{d}$ -distance between the process and the estimated one. This is one of the most widely used metrics over stationary processes, its properties include that entropy is  $\bar{d}$ -continuous and the class of ergodic processes is  $\bar{d}$ -closed [7], [8], [9].

Ornstein and Weiss [8] proved that for stationary processes isomorphic to i.i.d. processes, the empirical distribution of the  $k(n)$ -length blocks is a strongly consistent estimator of the  $k(n)$ -length parts of the process in  $\bar{d}$ -distance if and only if  $k(n) \leq (\log n)/h$ , where  $h$  denotes the entropy of the process.

In this paper, we estimate the  $n$ -length part of a process  $X$  by a Markov process of order  $k(n)$ . The transition probabilities of this Markov estimator process are the empirical conditional probabilities. We assume that the process  $X$  is non-null, that is, the conditional probabilities of the symbols given the pasts are separated from zero, and that the continuity rate of the process  $X$  is summable. These conditions are usually assumed in this area [3], [4], [5], [6]. The summability of the continuity rate implies that the process is isomorphic to an i.i.d. process [1].

We obtain not only an asymptotic rate of convergence result but also an explicit bound on the probability that the  $\bar{d}$ -distance of the Markov estimator from the process  $X$  is greater than  $\varepsilon$ . For i.i.d. processes, our rate asymptotically almost attains the best rate possible. Our results rely upon those in [4], [5], [6]. We are not aware of prior works addressing specifically the rate of convergence of statistical estimation of processes in  $\bar{d}$ -distance.

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## II. FINITE SAMPLES

Let  $X = \{X_i, -\infty < i < +\infty\}$  be a stationary ergodic stochastic process with finite alphabet  $A$ . We write  $X_i^j = X_i, \dots, X_j$  and  $x_i^j = x_i, \dots, x_j \in A^{j-i+1}$  for  $j \geq i$ . For two strings  $x_1^j \in A^j$  and  $y_1^m \in A^m$ ,  $x_1^j y_1^m$  denotes their concatenation  $x_1, \dots, x_j, y_1, \dots, y_m \in A^{j+m}$ . Write

$$P(x_j^m) = \Pr\{X_j^m = x_j^m\}$$

and, if  $P(x_{-m}^{-1}) > 0$ ,

$$P(a|x_{-m}^{-1}) = \Pr\{X_0 = a \mid X_{-m}^{-1} = x_{-m}^{-1}\}.$$

The process  $X$  is called *non-null* if always  $P(x_{-m}^{-1}) > 0$  and, in addition,

$$p_{\text{inf}} = \inf_{m \geq 1} \min_{a \in A, x_{-m}^{-1} \in A^m} P(a|x_{-m}^{-1}) > 0.$$

The *continuity rate* of the process  $X$  is

$$\gamma(k) = \sup_{m \geq k} \max_{a \in A} \max_{x_{-m}^{-1}, y_{-m}^{-1} \in A^m: x_{-k}^{-1} = y_{-k}^{-1}} |P(a|x_{-m}^{-1}) - P(a|y_{-m}^{-1})|.$$

Let  $\gamma = \sum_{k=1}^{\infty} \gamma(k)$  and

$$\alpha = \frac{1}{\prod_{j=1}^{+\infty} (1 - \gamma(j))}$$

and

$$\beta(k) = \frac{1 - (1 - |A|\gamma(k))^k}{k\gamma(k) \prod_{j=1}^{+\infty} (1 - |A|\gamma(j))^2}.$$

If  $\gamma < +\infty$ , the process  $X$  is said to have summable continuity rate. In this case,  $\alpha < +\infty$  and  $\beta(k) \leq \beta$  for some constant  $0 < \beta < +\infty$ .

The per-letter Hamming distance between two strings  $x_1^n$  and  $y_1^n$  is

$$d_n(x_1^n, y_1^n) = \frac{1}{n} \sum_{i=1}^n \mathbb{I}(x_i \neq y_i),$$

where

$$\mathbb{I}(a \neq b) = \begin{cases} 1 & \text{if } a \neq b \\ 0 & \text{if } a = b. \end{cases}$$

The  $\bar{d}$ -distance between two random sequences  $X_1^n$  and  $Y_1^n$  with distributions  $P_X$  and  $P_Y$ , respectively, is defined by

$$\bar{d}(X_1^n, Y_1^n) = \min_{\mathbb{P}} \mathbb{E}_{\mathbb{P}} d_n(\tilde{X}_1^n, \tilde{Y}_1^n),$$

where the minimum is taken over all the joint distributions  $\mathbb{P}$  of  $\tilde{X}_1^n$  and  $\tilde{Y}_1^n$  whose marginals are equal to  $P_X$  and  $P_Y$ . The *Kullback–Leibler distance* of these sequences is

$$D(X_1^n \| Y_1^n) = \sum_{a_1^n \in A^n} P_X(a_1^n) \frac{P_X(a_1^n)}{P_Y(a_1^n)}.$$

The process  $X$  is a *Markov chain* of order  $k$  if for each  $n > k$  and  $x_1^n \in A^n$

$$P(x_1^n) = P(x_1^k) \prod_{i=k+1}^n P(x_i | x_{i-k}^{i-1}),$$

where  $P(x_1^k)$  is called initial distribution and  $P(\cdot|\cdot)$  is called transition probability matrix. The case  $k = 0$  corresponds to an i.i.d. process. The  $k$ -order Markov approximation of a process  $X$  is the Markov chain, denoted by  $X[k]$ , of order  $k$  whose transition probabilities are  $P(a|a_1^k)$ ,  $a \in A$ ,  $a_1^k \in A^k$ .

Let  $N_n(a_1^k)$  denote the number of occurrences of the string  $a_1^k$  in the sample  $X_1^n$

$$N_n(a_1^k) = |\{i : X_{i+1}^{i+k} = a_1^k, 0 \leq i \leq n-k\}|.$$

The empirical probability of the string  $a_1^k$  is

$$\hat{P}_n(a_1^k) = \frac{1}{n-k+1} N_n(a_1^k).$$

The *empirical  $k$ -order Markov approximation* of a process  $X$  based on the sample  $X_1^n$  is the stationary Markov chain, denoted by  $\hat{X}[k]$ , of order  $k$  whose transition probabilities are the empirical conditional probabilities

$$\hat{P}_n(a|a_1^k) = \frac{N_n(a_1^k a)}{N_{n-1}(a_1^k)}, \quad a \in A, a_1^k \in A^k.$$

If the initial distribution of a stationary Markov chain with these transition probabilities is not unique, then any of these initial distributions can be taken.

*Theorem 1:* Let  $X$  be a non-null stationary ergodic process with summable continuity rate. Then, for any  $\varepsilon > 0$ , the empirical  $k$ -order Markov approximation of the process satisfies

$$\begin{aligned} & \Pr \left\{ \bar{d} \left( X_1^n, \hat{X}[k]_1^n \right) > \varepsilon \right\} \\ & \leq 2e^{1/e} |A|^{k+2} \exp \left\{ - \frac{(n-k) p_{\text{inf}}^{2k+2}}{16e|A|^3 (|A|\gamma + p_{\text{inf}})(k+1)} \right. \\ & \quad \left. \left[ \left( \frac{\varepsilon - \beta(k) p_{\text{inf}}^{-2} \gamma(k)}{\alpha + 1} \right)^2 - \frac{k |\log p_{\text{inf}}|}{2n} \right] \right\}. \end{aligned}$$

*Proof:* The proof relies upon the results in [4], [5], [6]. See [2] for the details. ■

### III. ASYMPTOTIC RESULTS

Based on the results for finite samples of size  $n$ , we will derive asymptotic bounds as  $n \rightarrow \infty$ .

*Theorem 2:* Let  $X$  be a non-null stationary ergodic process with summable continuity rate. Then, for any  $\mu > 0$ , the empirical  $(\nu \log n)$ -order Markov approximation of the process satisfies

$$\bar{d} \left( X_1^n, \hat{X}[\nu \log n]_1^n \right) \leq \frac{\beta(\nu \log n)}{p_{\text{inf}}^2} \gamma(\nu \log n) + \frac{1}{n^{1/2-\mu}}$$

eventually almost surely as  $n \rightarrow \infty$ , if

$$\nu < \frac{\mu}{|\log p_{\text{inf}}|}.$$

*Proof:* For  $\varepsilon = \beta(\nu \log n) p_{\text{inf}}^{-2} \gamma(\nu \log n) + 1/(n^{1/2-\mu})$  and  $k = \nu \log n$ , Theorem 1 yields

$$\begin{aligned} & \Pr \left\{ \bar{d} \left( X_1^n, \hat{X}[\nu \log n]_1^n \right) > \right. \\ & \quad \left. \frac{\beta(\nu \log n)}{p_{\text{inf}}^2} \gamma(\nu \log n) + \frac{1}{n^{1/2-\mu}} \right\} \\ & \leq 2e^{1/e} |A|^{2+\nu \log n} \\ & \quad \exp \left\{ - \frac{p_{\text{inf}}^2}{16e|A|^3 (|A|\gamma + p_{\text{inf}})(\alpha + 1)^2} \right. \\ & \quad \left. \frac{(n - \nu \log n) n^{-2\nu |\log p_{\text{inf}}|}}{(1 + \nu \log n)n} \right. \\ & \quad \left. \left[ n^{2\mu} - \frac{\nu |\log p_{\text{inf}}| (\alpha + 1)^2 \log n}{2} \right] \right\}, \end{aligned}$$

which is summable in  $n$  and allows application of the Borel–Cantelli lemma. ■

*Remark 1:* While the parameter  $\nu$  in Theorem 2 depends on the actual process  $X$ , it holds for all  $X$  satisfying the hypothesis that any  $k_n \rightarrow +\infty$  with  $k_n = o(\log n)$  guarantees  $\bar{d} \left( X_1^n, \hat{X}[k_n]_1^n \right) \rightarrow 0$ , almost surely.

*Remark 2:* In Theorem 2, in the upper bound the first term is the bias due to the error of the approximation of the process by a Markov chain. The second term is the variation due to the error of the estimation of the parameters of the Markov chain based on a sample.

*Remark 3:* If the process is i.i.d., then  $\bar{d} \left( X_1^n, \hat{X}[0]_1^n \right) = \bar{d} \left( X[0]_1^n, \hat{X}[0]_1^n \right)$  equals the variation distance [9], whose order is  $n^{-1/2}$ . Since in this case  $\gamma(k) = 0$ ,  $k = 0, 1, \dots$ , the order of the upper bound in Theorem 2 cannot be improved significantly.

*Corollary 1:* Let  $X$  be a non-null stationary ergodic process with continuity rate  $\gamma(k) = c' \exp(-ck)$ ,  $k = 1, 2, \dots$ , where  $c', c > 0$  are constants. Then, for any  $\mu > 0$ , the empirical  $(\nu \log n)$ -order Markov approximation of the process satisfies

$$\bar{d} \left( X_1^n, \hat{X}[\nu \log n]_1^n \right) \leq \frac{2}{n^{1/2-\mu}}$$

eventually almost surely as  $n \rightarrow \infty$ , if

$$\nu < \frac{\mu}{|\log p_{\text{inf}}|}$$

and

$$c' \leq \frac{p_{\text{inf}}^2}{\beta} \quad \text{and} \quad c \geq \frac{1}{\nu} \left( \frac{1}{2} - \mu \right).$$

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